

Risk Transmission to Fintech Firms: Is WAR Different From COVID?

Farrukh Nawaz, 1 Fakhrul Hasan, 2 Tonmoy Choudhury, 3 and Umar Kayani 4

¹Faculty of Business Studies, Arab Open University, Riyadh, Saudi Arabia ²Newcastle Business School, Northumbria University, Newcastle, UK

³Rochester Institute of Technology (RIT), Dubai, United Arab Emirates

⁴School of Business & Quality Management, Hamdan Bin Mohammed Smart University, Dubai, United Arab Emirates

(Received 13 May 2025; Revised 11 August 2025; Accepted 29 September 2025; Published online 31 October 2025)

Abstract: This study's main aim is to determine whether either the VIX (CBOE Volatility Index) in COVID-19 or the WAR, a reference to the Russian Invasion, had a more significant (negative) impact on Fintech firms' returns. The primary research method is the DCC-GARCH model, which is used to determine the conditional correlation between Fintech firms' returns and the VIX/OVX. We used nine Fintech firms from four different countries from December 13, 2019, to October 06, 2022. Our findings show that the VIX has an inverse impact on the returns of the Fintech sample in our study. For COVID, the war correlation remains above throughout the first half, while the situation changes during the second. We can also observe a drop in the correlation for COVID around day 30.

Keywords: COVID; stock market; return; volatility; WAR

I. INTRODUCTION

Fintech is the best example of how different kinds of shocks can influence technology adoption in unexpected ways in the post-COVID world [1,2]. It is defined as financial services innovation that is made possible and that results in new business models, applications, procedures, or goods that have a significant effect on financial markets, financial institutions, and financial services. In addition, these structural shocks can potentially alter communities and economies in the long run, as proven by numerous authors [3,4]. Recently, the devastating impact of COVID and its rapid transmission rate have forced communities to enact lockdown procedures and/or self-imposed behavioral adjustments [5–14]. Despite the high monetary and human costs, there have been some winners and potential benefits of these measures. There is a growing anecdotal evidence that the adoption and utilization of technology, particularly that enables businesses to communicate and exchange goods and services over distance, have significantly increased due to COVID [15-20]. In fact, it is suggested that the use of digital banking has assisted many people and businesses in reducing some of the health risks and unfavorable socioeconomic repercussions of the epidemic. This rapid adoption will probably have a significant impact on the market equilibrium between established incumbents and emerging tech-based firms in financial intermediation risk, which can be calculated by the Chicago Board Options Exchange's CBOE Volatility Index, or the VIX.

This study's main aim is to determine whether either the VIX in COVID-19 or the WAR (a reference to the Russian Invasion) had a more significant (negative) impact on Fintech firms' returns. To answer this question, this research paper uses two different research methods. The primary research method is the

DCC-GARCH model, which is used to determine the conditional correlation between Fintech firms' returns and the VIX/OVX (the daily log-returns) to determine whether the VIX/OVX returns Granger-Causes firms' stock returns. We also use positive and negative changes in these volatility indices, where a positive difference in the VIX/OVX index is VIX/OVX_U, and an adverse change in the VIX/OVX index is VIX/OVX_D. We used nine Fintech firms from four different countries (one from Australia, one from Hong Kong, one from the Netherlands, and six from the US) from December 13, 2019, to October 06, 2022. The nine Fintech firms are: (1) Visa Inc., (2) Tencent Holdings Ltd., (3) Mastercard Inc., (4) Intuit Inc., (5) PayPal Holdings Inc., (6) Fiserv Inc., (7) Adyen NV, (8) Bill.com Holdings Inc., and (9) Xero Ltd.

This work makes several contributions to the literature. First, this is the first time anyone has answered the question of whether the VIX in COVID-19 and WAR had any impact on Fintech firms' returns. Second, this is the first time anyone has used the DCC-GARCH model. Third, we use a more extensive data set than studies in this area.

Our findings show that the VIX had an inverse impact on the returns of the Fintech sample in our study. For COVID, the war correlation remains above throughout the first half, while the situation changes during the second. We can also observe a drop in the correlation for COVID around day 30. We can conclude that the impact pattern is unchanged during the crisis for the Fintech firms from VIX. The stock market and other markets crashed due to COVID-19 and the WAR in a short amount of time. As a result, the level of panic during COVID-19 and the WAR was significantly higher both in the real world and in the capital markets, which could account for the stronger correlation between increases in the VIX and Fintech firms.

The remainder of the paper is structured as follows. Section II discusses methodology. Our sample details and preliminary analysis are discussed in Section III. We report our results in Section IV, and in Section V, we conclude our paper.

II. LITERATURE REVIEW

The convergence of war and a global pandemic, such as the COVID-19 crisis, has presented an unprecedented challenge for various industries worldwide. Among these industries, financial technology (Fintech) firms, which have been rapidly evolving in recent years, are particularly sensitive to the disruptions caused by geopolitical conflicts and health crises. Ling and Pei [21] studied the impact of COVID-19 and its association with Fintech. The study adopted the quarterly data from Chinese publicly listed companies spanning from 2011 to 2020. Our findings indicate that the pandemic has amplified financial constraints for enterprises; however, the advancement of Fintech can alleviate its adverse impact. Moreover, Tripalupi and Anggahegari [22] studied the challenges and prospects of Shariah-based Fintech in Indonesia in light of the COVID-19 pandemic. Tripalupi and Anggahegari [22] studied primary focus on Shariah Fintech, the pandemic's influence on Fintech, and the opportunities and obstacles for Shariah Fintech. The study adopted a qualitative and descriptive approach for analysis. The findings highlight that by the close of 2019, Syariah Fintech assets had surged by 2074.81% compared to the preceding year. Despite pandemic repercussions, Fintech growth demonstrated relative stability during the COVID-19 crisis.

In the face of unparalleled challenges prompted by the advent of the COVID-19 pandemic, the embrace of Fintech solutions emerged as a pivotal strategy in addressing the evolving needs of individuals, businesses, and institutions [23-26]. The global lockdown due to the COVID-19 outbreak, plays a significant role in the adoption of Fintech due to its potential such as digital payments, remote account management, and contactless transactions. Fu and Mishra [27] examined the adoption of Fintech and digital finance by businesses due to the COVID-19 pandemic. The results revealed a significant increase in the rate of finance app downloads due to the spread of COVID-19 and governmentimposed lockdowns. Moreover, the landscape of financial transactions has undergone a significant transformation in the wake of the COVID-19 pandemic. This global crisis has propelled a notable shift in behavior, with individuals and businesses alike increasingly embracing Fintech solutions as a preferred avenue for their financial needs and services. The inherent advantages of Fintech, such as contactless transactions, remote account management, and digital payment options, have not only provided convenience but also aligned well with the health and safety concerns posed by the pandemic.

As another example of a shock, most financial market participants and Ukrainian citizens were caught off guard by Russia's invasion of their country, which was an exogenous shock for foreign companies with Russian operations in the financial market [28]. After military bases were built close to the border, Russia officially attacked Ukraine on February 24, 2022, converting the world's financial markets into a risk magnet. The ongoing Russia-Ukraine conflict has been shown to exert notable influences on various industries, and the Fintech sector is no exception. The evidence shows that the impacts of conflicts on Fintech firms shed light on the intricate interplay between global geopolitical dynamics and the rapidly evolving Fintech landscape. Allinger and Barisitz [29] studied the Russian Fintech system in terms of sanctions and war. The study revealed that the landscape is shaped by international sanctions. The challenges posed by sanctions affect major Fintech entities, influencing their advancement in the realm of digital ecosystems.

Various studies have adopted machine learning methods to study volatility behavior [30–35]. Furthermore, Michańków, Kwiatkowski [36] develops a hybrid approach to forecasting the volatility and risk of financial instruments by combining GARCH and neural network models. Four GARCH specifications are used: standard GARCH, EGARCH, GJR-GARCH, and APARCH in the research study. The models are tested using daily logarithmic returns on the S&P 500 index, gold prices, and Bitcoin prices, which represent distinct volatility dynamics. The results show that hybrid solutions give more accurate point volatility forecasts, but not necessarily better VaR and ES forecasts.

Wang and Liu [37] presents a multivariate volatility modeling framework called long short-term memory-enhanced BEKK (LSTM-BEKK), which addresses the limits of traditional multivariate GARCH models in dealing with persistent volatility clustering and asymmetric co-movements across assets. The empirical results from equity markets show that LSTM-BEKK provides better out-of-sample portfolio risk forecasts while keeping the interpretability of BEKK models.

Manogna and Dharmaji [38] examined the agricultural price volatility forecasting using a hybrid long short-term memory (LSTM)–GARCH model by adopting the historical data for 23 commodities across 165 markets in India from February 2010 to June 2024. The results indicate that the hybrid approach effectively addresses price instability and improves predictive performance.

Abakah and Adeabah [39] adopted an innovative methodology to study the public sentiments in time of the Russia-Ukraine conflict outbreak, investigating their implications for the global Fintech sector and blockchain markets. Abakah and Adeabah [39] constructed an index for analysis known as the Russia-Ukraine War Economic Sanctions News Sentiment Index (RUWESsent), incorporating media coverage, panic, sentiment, and media hype from sources such as Twitter sentiments, Google Trend, Wikipedia Trend, and News Sentiments. The finding of the study shows that the positive RUWESsent values have positively influenced the returns of Fintech and blockchain market stocks in bullish market conditions, while exerting a negative influence during bearish market periods. Notably, a heightened negative correlation is observed immediately following the invasion, revealing that adverse sentiments adversely impact Fintech and blockchain market stocks.

III. METHODOLOGY

In this research paper, we use the DCC-GARCH model to determine the conditional correlation between the firm's returns and the VIX/OVX (the daily log-returns) and to determine whether VIX/OVX returns Granger-Causes firms' stock returns. To conduct this research, we developed a bivariate (2×1) vector $\mathbf{R}_t = (R_{s,t}, R_{VIX,OVX,t})'$ of conditional stock returns of Fintech firm s, $R_{s,t}$, and VIX/OVX returns, $R_{VIX,OVX,t}$, using a DCC-AR(n)-GARCH (p,q) model following Engle [40]:

$$\mathbf{R}_t = \vartheta_t + \mathbf{\Lambda}_t^{0.5} \mathbf{\Theta}_t \tag{1}$$

where Θ_t is a two-by-one vector of independent and identically distributed innovations, and it is assumed that it has a bivariate normal distribution (that is, Θ_t has the form $\Theta_t \sim N(0, \Lambda_t)$). The conditional covariance matrix, denoted by " Λ_t ," is capable of being decomposed into the diagonal matrix D_t , which has dimensions of

4.9%

2.5%

Information

Australia

NZXROE0001S2

-0.8%

11.7

98.8%

-1.0%

6,742,568,639

Xero Ltd

XR0.AX

America

| Fechnology

United States of

US0900431000

-50.8%

3.1

43.8%

~6.6~

12,409,951,993

Bill.com Holdings

BILL.K

Technology

7.3%

Information Technology Information

Netherlands

NL0012969182

7.8%

18.9

7.9%

30.9%

38,871,586,535

Adyen NV

ADYEN.AS

WACC Equity Risk Premium 6.1% 4.9% 4.9% 4.9% 4.9% 4.9% Communication GICS Sector Name Information Information Information Information Technology Information Technology **Technology Technology Fechnology** Services Country of Exchange Hong Kong States of States of States of States of States of America America America America America United United United United United KYG875721634 US57636Q1040 US70450Y1038 US92826C8394 US3377381088 US4612021034 NIS! Tax Margin Income Aft 51.1% 40.7% 16.2% 16.4% 46.0% 8.0% 8 Price To Book Value Per Share 10.7 45.3 2.9 6.5 4.7 1.9 Total Debt to Total Equity 190.1% 42.1% 55.8% 41.6% 68.6% 40.1% Return on equity (%) 42.4% 153.8% 22.3% 15.7% 10.1% 6.4% 377,809,871,017 302,596,822,743 278,983,108,803 107,141,559,707 Market Cap 93,061,613,581 60,210,470,521 (OSD) Mastercard Fisery Inc Name Holdings Intuit Inc Holdings Tencent Visa Inc PayPal Сtd Identifier PYPL.0Q FISV.OQ 0700.HK INTU:0 MA.N (RIC) >

Sample information

Table I.

(Ahead of Print)

Table II. Descriptive statistics

	VIX	OVX	XERO	
Mean	0.001	0.001	0.000	
Variance	0.007	0.007	0.001	
Kurtosis	7.803	28.040	6.550	
Skewness	1.283	1.943	-0.799	
Jarque-Bera	908.157	19664.686	464.279	
<i>p</i> -value	0.000	0.000	0.000	
Augmented Dickey-Fuller (ADF)	-30.199	-27.844	-29.830	
<i>p</i> -value	0.000	0.000	0.000	
Count	735.000	735.000	735.000	
	INTUIT	VISA	MASTERCARD	TENCENT
Mean	0.001	0.000	0.000	0.000
Variance	0.001	0.000	0.001	0.001
Kurtosis	9.019	11.164	9.927	10.138
Skewness	0.047	-0.014	0.131	0.644
Jarque-Bera	1109.666	2041.196	1471.668	1611.183
<i>p</i> -value	0.000	0.000	0.000	0.000
Augmented Dickey-Fuller (ADF)	-31.585	-31.873	-29.647	-26.910
<i>p</i> -value	0.000	0.000	0.000	0.000
Count	735.000	735.000	735.000	735.000
	PAYPAL	FISERV	ADYEN	BILL
Mean	0.000	0.000	0.001	0.002
Variance	0.001	0.000	0.001	0.003
Kurtosis	14.276	12.894	4.640	8.069
Skewness	-0.922	-0.769	-0.067	0.780
Jarque-Bera	3998.250	3070.374	82.871	861.562
<i>p</i> -value	0.000	0.000	0.000	0.000
Augmented Dickey-Fuller (ADF)	-29.559	-32.264	-26.289	-26.035
<i>p</i> -value	0.000	0.000	0.000	0.000
Count	735.000	735.000	735.000	735.000

two by two and is made of conditional standard deviations $h_{i,t}^{0.5}$ with $i \in \{s, VIX\}$, and a conditional correlation matrix C_t :

$$\Lambda_t = D_t C_t D_t \tag{2}$$

Based on our assumption, we determine our conditional mean $-\vartheta_{i,t} = E(R_{i,t}|\mathcal{F}_{t-1})$ — is stated as a four-order autoregressive process, and the equation for our univariate conditional mean is as follows:

$$R_{i,t} = \vartheta_i + \sum_{j=1}^n \rho_j R_{i,t-j} + \sum_{p=1}^n \gamma_p R_{VIX,OVX,t-p} + \varepsilon_{i,t}$$
 (3)

where $\varepsilon_{i,t} = h_{i,t}^{0.5} z_t$ shows the innovations that were not standardized and had a mean of zero and a conditional variance of $h_{i,t}$. We used an AR term $\sum_{j=1}^{n} \rho_j R_{i,t-j}$ up to the order of n to consider the possibility of serial correlation in the returns of Fintech firms. Using the coefficient, we examine the Granger causation between the t-1lag in VIX/OVX returns and the t-1lag in the stock returns of Fintech businesses.

We followed Bollerslev [41] and measured the univariate conditional variances of stocks and VIX/OVX returns using the GARCH approach using $h_{i,t} = Var(R_{i,t}|\mathcal{F}_{t-1})$:

$$h_{i,t} = \varphi_{i,0} + \varphi_{i,1}e_{i,t-1}^2 + \sum_{i=1}^{3} \varphi_{i,2+j-1}h_{i,t-j}$$
 (4)

It is expected that every value of φ will be greater than zero. Our conditional correlation matrix can be represented by equation (5), using \mathbb{Q}_t (2 × 2) covariance matrix:

$$C_t = (\operatorname{diag} \mathbb{Q}_t)^{-0.5} \mathbb{Q}_t (\operatorname{diag} \mathbb{Q}_t)^{-0.5}$$
 (5)

According to Batten and Kinateder [42], if we specify u_{t-1} as (2×1) vector of standardized innovations with unconditional correlation matrix $\bar{\boldsymbol{C}}$, and $u_{i,t} = \frac{e_{i,t}}{\sqrt{h_{i,t}}}$, we can derive the matrix \mathbb{Q}_t as

$$\mathbb{Q}_{t} = (1 - \boldsymbol{a} - \boldsymbol{b})\bar{\boldsymbol{C}} + a\boldsymbol{u}_{t-1}\boldsymbol{u}\,\boldsymbol{\prime}_{t-1} + \boldsymbol{b}\mathbb{Q}_{t-1} \tag{6}$$

Positive scalers a and b's values are limited to a+b < 1. We can achieve conditional correlations from the components of the matrix \mathbb{Q}_t using Eq. (7), where $q_{s,VIX,t}$ is the conditional covariance $q_{s,VIX,t} = Cov(R_{s,t},R_{VIX,t}|\mathcal{F}_{t-1})$ between Fintech firm s and the VIX/OVX and $q_{s,s,t} = Var(R_{s,t}|\mathcal{F}_{t-1})$ and $q_{VIX,VIX,t} = Var(R_{VIX,t}|\mathcal{F}_{t-1})$ represent the conditional variances of Fintech firm s, and VIX/OVX returns, respectively.

$$\rho_{s,VIX,t} = q_{s,VIX,t} / (q_{s,s,t}q_{VIX,VIX,OVX,OVX,t})^{0.5}$$
(7)

We may use these estimates to compute the DCC-AR (n)-GARCH (p,q) model for our baseline result.

Table III. Pearson correlation

	VIX	OVX	VISA	MASTERCARD	TENCENT	INTUIT	PAYPAL	FISERV	ADYEN	BILL	XERO
VIX	1.000										
OVX	0.381	1.000									
VISA	-0.586	-0.305	1.000								
MASTERCARD	-0.600	-0.339	0.923	1.000							
TENCENT	-0.106	-0.130	0.154	0.165	1.000						
INTUIT	-0.577	-0.258	0.673	0.684	0.189	1.000					
PAYPAL	-0.504	-0.224	0.565	0.566	0.119	0.680	1.000				
FISERV	-0.513	-0.297	0.794	0.779	0.161	0.625	0.539	1.000			
ADYEN	-0.335	-0.123	0.346	0.327	0.171	0.422	0.420	0.281	1.000		
BILL	-0.401	-0.146	0.394	0.402	0.173	0.537	0.537	0.347	0.383	1.000	
XERO	-0.128	-0.139	0.249	0.231	0.311	0.260	0.201	0.275	0.306	0.169	1.000

A. DATA AND PRELIMINARY ANALYSIS

Daily closing price date, $P_{i,t}$, were collected from DataStream, consisting of nine Fintech firms from four different countries (one from Australia, one from Hong Kong, one from the Netherlands, and six from the US) for the period from December 13, 2019, to October 06, 2022. We chose this data frequency because it enabled us to simulate conditional correlations during a crisis. The nine Fintech firms are: (1) Visa Inc., (2) Tencent Holdings Ltd., (3) Mastercard Inc., (4) Intuit Inc., (5) PayPal Holdings Inc., (6) Fiserv Inc., (7) Adyen NV, (8) Bill.com Holdings Inc., and (9) Xero Ltd. Table I reports on the sample information.

The descriptive statistics are shown in Table II. From there, we can see that the mean value for both VIX and OVX is 0.001, and for both VIX and OVX, the skewness and kurtosis values are positive.

The descriptive statistics for the study's whole sample are shown in this table, encompassing both COVID-19 and WAR. The variables are (with RIC code in parentheses) the daily log-returns of CBOE Volatility Index-VIX, CBOE Crude Oil Volatility Index-OVX, VISA (V), MASTERCARD (MA.N), TENCENT (0700.HK), INTUIT (INTU.O), PAYPAL (PYPL.OQ), FISERV (FISV.OQ), ADYEN (ADYEN.AS), BILL (BILL.K), and XERO (XRO.AX). The sample period for all the variables is from December 13, 2019, to October 06, 2022.

Table IV. Baseline DCC-AR (n)-GARCH (p,q) Model Results between VIX and Fintech firms

	VISA	MASTERCARD	TENCENT	INTUIT	PAYPAL	FISERV	ADYEN	BILL	XERO	
Panel A: Mean Equation										
μ	0.001	0.001	0.000	0.002***	0.001	0.000	0.002*	0.003**	0.001*	
<i>p</i> -value	0.316	0.140	0.840	0.000	0.166	0.951	0.056	0.048	0.083	
$ ho_1$	-0.108**	-0.063	-0.040	-0.050	-0.111**	-0.095**	-0.017	0.073**	-0.107***	
<i>p</i> -value	0.026**	0.189**	0.285**	0.228***	0.012***	0.033***	0.683**	0.050	0.003	
λ_1	-0.022	-0.018	-0.046	0.001	0.002	-0.007	-0.042	0.055***	-0.096***	
<i>p</i> -value	0.025	0.103	0.000	0.936	0.902	0.448	0.001	0.006	0.000	
	Panel B: Variance Equation									
w_0	0.000***	0.000***	0.000**	0.000**	0.000***	0.000***	0.000**	0.000	0.000**	
<i>p</i> -value	0.000	0.000	0.030	0.039	0.009	0.001	0.036	0.126	0.044	
w_1	0.119***	0.149***	0.124***	0.100**	0.204***	0.099***	0.068***	0.055***	0.114***	
<i>p</i> -value	0.000	0.000	0.003	0.018	0.000	0.000	0.000	0.002	0.001	
w	0.829***	0.797***	0.754***	0.833***	0.795***	0.840***	0.917***	0.938***	0.836***	
<i>p</i> -value	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	
			Par	nel C: DCC E	Equation					
a	0.031***	0.035***	0.014	0.039	0.001	0.080*	0.100**	0.003	0.022	
<i>p</i> -value	0.004	0.003	0.272	0.121	0.782	0.056	0.038	0.525	0.584	
b	0.941***	0.945***	0.961***	0.809***	0.987***	0.396	0.137	0.991***	0.000	
<i>p</i> -value	0.000	0.000	0.000	0.000	0.000	0.107	0.557	0.000	1.000	
Panel D: Diagnostics Tests										
Log-likelihood	2912.862	2827.951	2610.078	2767.187	2488.749	2802.968	2476.529	2206.035	2546.63	
AIC	-5771.723	-5601.902	-5164.156	-5472.375	-4923.498	-5551.936	-4899.058	-4356.07	-5061.26	
BIC	-5647.674	-5477.853	-5035.512	-5329.948	-4799.449	-5427.886	-4775.009	-4227.427	-4987.684	
d-Stat	1.99493	1.987158	1.942384	1.991951	1.983993	2.006851	1.967216	1.960112	2.001641	

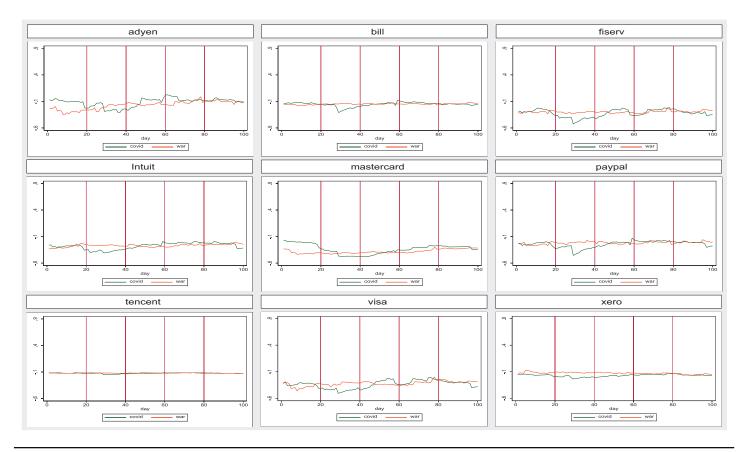


Fig. 1. COVID-19 vs. WAR Sample Correlation between VIX and Fintech firms during 100 significant days.

Table V. Baseline DCC-AR (n)-GARCH (p,q) Model Results between OVX and Fintech firms

	VISA	MASTERCARD	TENCENT	INTUIT	PAYPAL	FISERV	ADYEN	BILL	XERO
Panel A: Mean Equation									
μ	0.003	-0.004	-0.003	-0.004	-0.004	0.002	0.004	0.005	0.002
p-value	0.230	0.059	0.144	0.045	0.040	0.369	0.118	0.049	0.418
$ ho_1$	-0.086*	-0.078**	-0.042	-0.060	-0.058	-0.084*	-0.115**	-0.090*	-0.091*
p-value	0.066	0.041	0.260	0.107	0.112	0.080	0.016	0.059	0.056
λ_1	-0.544***	-0.277***	-0.114	-0.151*	-0.114*	-0.401***	-0.249***	-0.078	-0.122
p-value	0.000	0.005	0.156	0.059	0.062	0.004	0.003	0.105	0.155
Panel B: Variance Equation									
w_0	0.001***	0.001***	0.001***	0.001***	0.001***	0.001***	0.001***	0.001***	0.001***
p-value	0.000	0.007	0.008	0.003	0.005	0.000	0.000	0.000	0.000
w_1	0.258***	0.131***	0.132***	0.119***	0.122***	0.271***	0.269***	0.260***	0.366***
p-value	0.000	0.001	0.001	0.000	0.001	0.000	0.000	0.000	0.000
w_2	0.668***	0.785***	0.794***	0.797***	0.806***	0.665***	0.671***	0.672***	0.770***
p-value	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
			Pa	nel C: DCC E	Equation				
a	0.028	0.015	0.003	0.020	0.025	0.027	0.031	0.010	0.012
p-value	0.150	0.202	0.859	0.417	0.418	0.183	0.122	0.606	0.547
b	0.871***	0.969***	0.878***	0.867***	0.820***	0.846***	0.929***	0.838***	0.889***
p-value	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Panel D: Diagnostics Tests									
Log-likelihood	2837.644	2889.72	2708.852	2799.745	2622.07	2775.771	2486.364	2084.349	2484.695
AIC	-5621.287	-5723.44	-5361.704	-5543.49	-5188.139	-5497.542	-4918.729	-4114.698	-4907.39
BIC	-5497.238	-5594.796	-5233.061	-5414.846	-5059.495	-5373.493	-4794.68	-3990.649	-4764.963
d-Stat	2.023	2.006	2.018	2.011	2.008	2.005	1.939	1.970	1.960

For each pair of variables utilized in the study, the pairwise correlation and corresponding p-values are shown in Table III. Table III shows that VIX and the OVX are negatively correlated with all Fintech firms. On the other hand, we can see that all Fintech firms are positively correlated with each other.

The sample period for all the variables is from December 13, 2019, to October 06, 2022. All the correlations are significant at a 5% significance level, but we are not reporting the significance level for brevity.

IV. RESULTS

A. DCC-AR(n)-GARCH (p,q) RESULTS - VIX

The results from the base DCC model are shown in Table IV. Based on the DCC-AR (n)-GARCH (p,q) model, Table IV shows the bivariate association between the variables in our sample. The estimated coefficients are reproduced in four (4) different panels. The results of the mean equation are shown in Panel A, along with the corresponding p-values. The outcomes according to the variance equation are shown in Panel B. The DCC equation is the base for Panel C's results. Finally, the results in Panel D are based on different diagnostics tests. In addition, Table IV reports highly significant results in parameters a and b in the DCC equation. The significance level is mostly 1% and 5%. The results reinforce the fact that the association between the VIX and the Fintech firms of our sample is highly time-varying, as proven by a and b coefficients of the DCC equation of panel c. Similarly, the w1 and w2 parameters are also highly significant at a 1% significance level. This overall significance confirms that the Granger causality exists. This guarantees that the jump in the CBOE Volatility Index-VIX has an inverse impact on the returns of the Fintech sample in our study. Finally, Panel D reports the results of the diagnostic tests. The results based on the (AIC), (BIC), (d-statistic), and Loglikelihood confirm the model fitness.

Figure 1 plots the time-varying pairwise conditional correlations between the next-trading-day log returns of the CBOE Volatility Index-VIX and the Fintech firms in our sample, that is, VISA (V), MASTERCARD (MA.N), TENCENT (0700. HK), INTUIT (INTU.O), PAYPAL (PYPL.OQ), FISERV (FISV.OQ), ADYEN (ADYEN.AS), BILL (BILL.K), and XERO (XRO.AX). The green line in all the plots represents the DCC-GARCH correlations during COVID-19, whereas the correlations during the WAR are illustrated with green lines. Altogether, we can observe a similar picture across the firms for both crises. This affirms a strong similarity in the impact of the risk for the sector. For COVID, the war correlation remains above throughout the first half, and the situation changes during the second. We can also observe a drop in the correlation for COVID at around day 30. We can conclude that the impact pattern is unchanged during the crisis for Fintech firms from the VIX.

B. DCC-AR(n)-GARCH (p,q) RESULTS - OVX

The daily log-returns of the OVX and the Fintech firms in our sample are reported in Table V. The results in Table V confirm the time-varying correlations in our study following the previous sections. Similarly, the Granger Casualty is also established in our relationship based on the results of λ_1 of the mean equation in panel a of Table V. Furthermore, the results of the variance equation based on w1 and w2 are also highly significant. The results in Table V overall confirm a negative relationship between

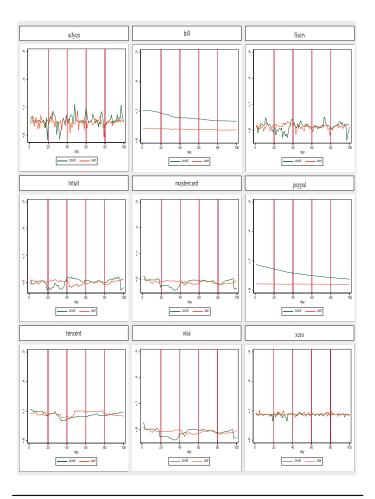


Fig. 2. COVID-19 vs. WAR Sample Correlation between OVX and Fintech firms over 100 significant days.

the OVX and the sample of Fintech firms. This shows that an increase in the OVX will have an inverse effect on Fintech firms. Finally, diagnostics tests were carried out to test the authenticity of the GARCH model and our results. The results confirm the fitness of the model used in this study.

Figure 2 plots the time-varying pairwise conditional correlations between the next-trading-day log returns, that is, $R_t = ln(P_t) - ln(P_{t-1})$ of OVX, and the Fintech firms in our sample, that is, VISA (V), MASTERCARD (MA.N), TENCENT (0700. HK), INTUIT (INTU.O), PAYPAL (PYPL.OQ), FISERV (FISV.OQ), ADYEN (ADYEN.AS), BILL (BILL.K), and XERO (XRO.AX). For both crises, the overall picture across the firms is mostly consistent, aside from some periodic spikes within the acceptable range. This proves that the impact of the oil risk on the sector is very similar to the VIX results. We can observe a slight fluctuation for prominent US players in the market, like Bill and PayPal. All correlations stay well below zero at all times.

Figure 3 plotted the time-varying pairwise conditional correlations between the next-trading-day log returns, that is, $R_t = ln(P_t) - ln(P_{t-1})$ VIX-OVX, and the Fintech firms in our sample, that is, VISA (V), MASTERCARD (MA.N), TENCENT (0700.HK), INTUIT (INTU.O), PAYPAL (PYPL.OQ), FISERV (FISV.OQ), ADYEN (ADYEN.AS), BILL (BILL.K), and XERO (XRO.AX). The selected period for COVID-19 was January 30, 2020, to June 17, 2020.

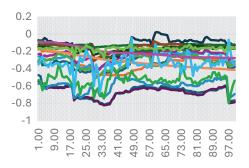


Fig. 3. DCC correlations between VIX and OVX with Fintech firms during 100 significant days of COVID-19.

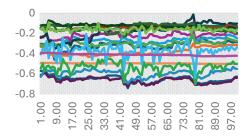


Fig. 4. DCC correlations between VIX and OVX with Fintech firms during 100 significant days of WAR.

Next, Figure 4 plots the time-varying pairwise conditional correlations between the next-trading-day log returns, that is, $R_t = ln(P_t) - ln(P_{t-1})$ of the VIX/OVX, and the Fintech firms in our sample, that is, VISA (V), MASTERCARD (MA.N), TENCENT (0700.HK), INTUIT (INTU.O), PAYPAL (PYPL.OQ), FISERV (FISV.OQ), ADYEN (ADYEN.AS), BILL (BILL.K), and XERO (XRO.AX). The selected period for WAR was February 24, 2022, to July 13, 2022. They cemented our previous findings for this purpose.

V. CONCLUSION

This paper explored how risk affected Fintech firms in the COVID-19 and the Russia-Ukraine war periods using nine Fintech firms from four different countries covering December 13, 2019, to October 06, 2022. Our main goal was to document whether the VIX in either COVID-19 or WAR had a more significant (negative) impact on Fintech firms' returns or not. In this regard, we could divide our findings into two subsections focusing on COVID-19 and WAR.

Our findings showed that the CBOE Volatility Index-VIX had an inverse impact on the returns of the Fintech sample in our study. We could also observe a drop in the correlation for COVID from around day 30.

Overall, these findings significantly impact the current literature on Fintech as they open the door for safe haven properties for Fintech stocks in the crisis period. However, as Fintech stocks are not traditional assets due to their connection to innovative design, we may need to look further into their capacity in the future.

CONFLICT OF INTEREST STATEMENT

The authors declare that they have no conflicts of interest to report regarding the present study.

REFERENCES

- [1] X. Xie and X. Zhu, "FinTech and capital allocation efficiency: another equity-efficiency dilemma?" *Glob. Finance J.* vol. 2022, Article ID 100741, 2022.
- [2] F. Naz, S. Karim, A. Houcine, and M. A. Naeem, "Fintech growth during COVID-19 in MENA region: current challenges and future prospects," *Electr. Commerce Res.* vol. 24, no. 1, pp. 371– 392, 2024.
- [3] J. A. Batten, T. Choudhury, H. Kinateder, and N. F. Wagner, "Volatility impacts on the European banking sector: GFC and COVID-19," Ann. Operat. Res., pp. 1–26, 2022.
- [4] M. K. Hassan, H. G. Djajadikerta, T. Choudhury, and M. Kamran, "Safe havens in Islamic financial markets: COVID-19 versus GFC," *Glob. Finance J.*, vol. 54, Article ID 100643, 2022.
- [5] Y. P. S. Balhara, D. Kattula, S. Singh, S. Chukkali, and R. Bhargava, "Impact of lockdown following COVID-19 on the gaming behavior of college students," *Indian J. Public Health*, vol. 64, no. 6, pp. 172– 176, 2020.
- [6] W.-C. Chin, T.-C. Yao, I. Tang, P.-Y. Lee, and Y.-S. Huang, "The impact of COVID-19 lockdown on sleep patterns, emotions, and behaviors of children and adolescents in Taiwan," *Front. Psychiatry*, vol. 13, Article ID 975399, 2022.
- [7] P. Dalal, D. Roy, P. Choudhary, S. K. Kar, and A. Tripathi, "Emerging mental health issues during the COVID-19 pandemic: An Indian perspective," *Indian J. Psychiatry*, vol. 62, Suppl. 3, pp. S354–S364, 2020.
- [8] M. W. Hussain, T. Mirza, and M. M. Hassan, "Impact of COVID-19 pandemic on the human behavior," *Int. J. Educ. Manag. Eng.*, vol. 10, no. 8, pp. 35–61, 2020.
- [9] B. Janz, E. Kaznina, C. Ammann, D. Kohlberg, and C. Mamali, "COVID-19 forced social distancing and isolation: a multi-perspective experience," *Ethics Prog.*, vol. 11, no. 1, pp. 20–60, 2020.
- [10] A. Kitkowska, A. S. Alaqra, and E. Wästlund, "Lockdown locomotion: the fast-forwarding effects of technology use on digital well-being due to COVID-19 restrictions," *Behav. Inform. Technol.*, vol. 43, no. 6, pp. 1178–1205, 2024.
- [11] S. R. Kraaijeveld, "COVID-19: against a lockdown approach," *Asian Bioethics Rev.*, vol. 13, no. 2, pp. 195–212, 2021.
- [12] P. K. Panda et al., "Psychological and behavioral impact of lockdown and quarantine measures for COVID-19 pandemic on children, adolescents and caregivers: a systematic review and meta-analysis," *J. Tropical Pediatr.*, vol. 67, no. 1. fmaa122, 2021.
- [13] K. Shulla et al., "Effects of COVID-19 on the sustainable development goals (SDGs)," *Discov. Sustain.*, vol. 2, pp. 1–19, 2021.
- [14] M. P. Tavolacci et al., "The impact of COVID-19 lockdown on health behaviors among students of a French university," *Int. J. Environ. Res. Public Health*, vol. 18, no. 8, pp. 4346, 2021.
- [15] N. Broberg and G. Golden, "How are OECD governments navigating the digital higher education landscape? Evidence from a comparative policy survey," OECD Educ. Work. Papers, vol. 303, pp. 1–54, 2023.
- [16] J. Huňady, P. Pisár, S. Khawaja, and F. H. Qureshi, "The digital transformation of european union countries before and during COVID-19," *Bus. Syst. Res.*, vol. 15, no. 1, pp. 22–45, 2024.
- [17] S. Khumalo, M. Mlotshwa, Z. Khumalo, and O. Raphalo, editors, "The impact of COVID-19 on e-commerce through a systematic review." In *European Conference on Innovation and Entrepreneur-ship*. Academic Conferences International Limited; 2023.
- [18] A. A. Lashitew, "When businesses go digital: the role of CEO attributes in technology adoption and utilization during the

- COVID-19 pandemic," *Technolog. Forecast. Social Change*, vol. 189, pp. 122324, 2023.
- [19] R. van der Vlies, "Digital strategies in education across OECD countries: Exploring education policies on digital technologies," OECD Education Working Papers, vol. 226, pp. 1–45, 2020.
- [20] T. Zou and A. Cheshmehzangi, "ICT adoption and booming ecommerce usage in the COVID-19 era," Front. Psychol., vol. 13, Article ID 916843, 2022.
- [21] S. Ling, T. Pei, Z. Li, and Z. Zhang, "Impact of COVID-19 on financial constraints and the moderating effect of financial technology," *Emerg. Mark. Finance Trade*, vol. 57, no. 6, pp. 1675–1688, 2021.
- [22] R. I. Tripalupi and P. Anggahegari, "The impact of Covid-19 pandemic: challenges and opportunities of Syariah Financial Technology", *Int. J. Nusantara Islam*, vol. 8, no. 1, pp. 119–128, 2020.
- [23] O. C. Edo et al., "Fintech adoption dynamics in a pandemic: an experience from some financial institutions in Nigeria during COVID-19 using machine learning approach," *Cogent Bus. Manag.*, vol. 10, no. 2, Article ID 2242985, 2023.
- [24] J. Fu and M. Mishra, "The global impact of COVID-19 on FinTech adoption," Swiss Finance Institute Research Paper 2020(20-38).
- [25] A. Khaliq, The impact of COVID-19 on FinTech lending in Indonesia: Evidence from interrupted time series analysis. ar-Xiv:250506655, 2025.
- [26] O. Shaikh, C. Ung, D. Yang, and F. A. Chacon, "Six feet apart: online payments during the COVID-19 pandemic," in *Proceedings of the ACM on Human-Computer Interaction*, vol. 6, no. CSCW2, pp. 1–33, 2022.
- [27] J. Fu and M. Mishra, Fintech in the time of COVID-19: Trust and technological adoption during crises. 2020.
- [28] S. Boubaker, J. W. Goodell, D. K. Pandey, and V. Kumari, "Heterogeneous impacts of wars on global equity markets: evidence from the invasion of Ukraine," *Finance Res. Lett.*, vol. 48, Article ID 102934, 2022
- [29] K. Allinger, S. Barisitz, and A. Timel, "Russia's large fintechs and digital ecosystems in the face of war and sanctions," *Focus Eur. Econ. Integr. Q.*, vol. 3, pp. 47–65, 2022.
- [30] A. Chatterjee, H. Bhowmick, J. Sen, editors, "Stock volatility prediction using time series and deep learning approach," in 2022 IEEE

- 2nd Mysore Sub Section International Conference (MysuruCon). IEEE; 2022.
- [31] V. D'Amato, S. Levantesi, and G. Piscopo, "Deep learning in predicting cryptocurrency volatility," *Physica A*, vol. 596, Article ID 127158, 2022.
- [32] S. M. Daniali et al., "Predicting volatility index according to technical index and economic indicators on the basis of deep learning algorithm," *Sustainability*, vol. 13, no. 24, Article ID 14011, 2021
- [33] Y. Liu, "Novel volatility forecasting using deep learning-long short term memory recurrent neural networks," *Exp. Syst. Appl.*, vol. 132, pp. 99–109, 2019.
- [34] M. Sardelich and S. Manandhar, Multimodal deep learning for shortterm stock volatility prediction. arXiv:181210479, 2018.
- [35] Y. Song, X. Tang, H. Wang, and Z. Ma, "Volatility forecasting for stock market incorporating macroeconomic variables based on GARCH-MIDAS and deep learning models," *J. Forecast.* vol. 42, nol. 1, pp. 51–59, 2023.
- [36] J. Michańków, Ł. Kwiatkowski, and J. Morajda, Combining Deep Learning and GARCH Models for Financial Volatility and Risk Forecasting. arXiv:231001063, 2023.
- [37] H. Wang, C. Liu, M.-N. Tran, and C. Wang, Deep Learning Enhanced Multivariate GARCH. arXiv:250602796, 2025.
- [38] R. Manogna, V. Dharmaji, and S. Sarang, "A novel hybrid neural network-based volatility forecasting of agricultural commodity prices: empirical evidence from India," *J. Big Data.*, vol. 12, no. 1, pp. 85, 2025.
- [39] E. J. A. Abakah, D. Adeabah, A. K. Tiwari, and M. Abdullah, Analyzing the Effect of Public Sentiment Towards Economic Sanctions News during Russia-Ukraine Conflict on Blockchain Market and Fintech Industry. Available at SSRN 4359071. 2022.
- [40] R. Engle, "Dynamic conditional correlation: a simple class of multivariate generalized autoregressive conditional heteroskedasticity models," *J. Bus. Econ. Stat.*, vol. 20, no. 3, pp. 339–350, 2002.
- [41] T. Bollerslev, "Generalized autoregressive conditional heteroskedasticity," *J. Econometr.*, vol. 31, no. 3, pp. 307–327, 1986.
- [42] J. A. Batten, H. Kinateder, P. G. Szilagyi, and N. F. Wagner, "Hedging stocks with oil," *Energy Econ.*, vol. 93, Article ID 104422, 2021.